# DETECTION THRESHOLD IN VERY SPARSE MATRIX COMPLETION

Charles Bordenave

Work in collaboration with Raj Rao Nadakuditi (Univ. Michigan) and Simon Coste (Inria Paris)

ArXiv:2005.06062

# MATRIX ESTIMATION

# MATRIX ESTIMATION

Let  $P \in M_{n,m}(\mathbb{R})$  be a large rectangular matrix  $n = \Theta(m)$ .

We observe each entry independently with probability d/n. The other entries remain hidden.

d = average number of observed entries per row.

We assume that the matrix P has a simple structure: notably small rank and some spectral incoherence.

# COMPLETION AND ESTIMATION

Matrix completion: can we reconstruct exactly P thanks to this observation?

Possible in the regime  $d \ge C \log n$ .

#### COMPLETION AND ESTIMATION

Matrix completion: can we reconstruct exactly P thanks to this observation?

Possible in the regime  $d \ge C \log n$ .

Matrix estimation: we look for a matrix with a small mean square error

$$MSE(\hat{P}) = \sum_{i,j} |\hat{P}_{ij} - P_{ij}|^2 = \|\hat{P} - P\|_F^2.$$

Best bounds for  $d \ll \log n$ :  $MSE(\hat{P}) = O(mn/d)$ .

Candès-Tao 09, Candès-Recht 10, Keshavan-Montanari-Oh 09 ....

#### PRINCIPAL COMPONENT ANALYSIS

Singular value decomposition of  $P \in M_{n,m}(\mathbb{C})$ :

$$P = UDV^* = \sum_{k=1}^n s_k u_k v_k^*,$$

where  $D = \text{diag}(s_1, \ldots, s_n) \in M_{n,m}(\mathbb{C})$  and  $s_1 \ge \ldots \ge s_n \ge 0$ are the singular values de P.



# MATRIX DETECTION

 $\star$  Above which value of d can we reconstruct a consistent estimator of  $s_k?$ 

# MATRIX DETECTION

\* Above which value of d can we reconstruct a consistent estimator of  $s_k$ ?

\* Fix  $0 < \gamma < 1$ . Find the smallest d such that there is with high probability an estimator  $\hat{u}_k$  of  $u_k$  with

 $|\langle \hat{u}_k, u_k \rangle| \geqslant \gamma.$ 

# MATRIX DETECTION

\* Above which value of d can we reconstruct a consistent estimator of  $s_k$ ?

\* Fix  $0 < \gamma < 1$ . Find the smallest d such that there is with high probability an estimator  $\hat{u}_k$  of  $u_k$  with

 $|\langle \hat{u}_k, u_k \rangle| \geqslant \gamma.$ 

\* Fix  $0 < \gamma < 1$ . Find the smallest d such that there is with high probability an estimator  $\hat{P}$  of P with

$$MSE(\hat{P}) = \|P - \hat{P}\|_F^2 \ge \gamma \|P\|_F^2.$$

# Applications

Numerous applications in global positioning, remote sensing, signal processing, computer vision, ... but the most famous is collaborative filtering.

# Applications

Numerous applications in global positioning, remote sensing, signal processing, computer vision, ... but the most famous is collaborative filtering.

*Guess what a user likes even before she knows it.* The Netflix prize launched in 2006 consisted in minimizing the MSE (on a sample) with respect to the matrix:

 $P_{ij} = \text{mark given by user } i \text{ on movie } j.$ 

Let  $P \in M_n(\mathbb{R})$  be a symmetric matrix.

Let  $M = (M_{ij}) \in M_n(\mathbb{R})$  with  $M_{ij} \in \{0, 1\}$  iid Bernoulli with parameter d/n.

Let  $P \in M_n(\mathbb{R})$  be a symmetric matrix.

Let  $M = (M_{ij}) \in M_n(\mathbb{R})$  with  $M_{ij} \in \{0, 1\}$  iid Bernoulli with parameter d/n.

The non-symmetric observed matrix  $A \in M_n(\mathbb{R})$  is

$$A_{ij} = \frac{n}{d} P_{ij} M_{ij}.$$

Let  $P \in M_n(\mathbb{R})$  be a symmetric matrix.

Let  $M = (M_{ij}) \in M_n(\mathbb{R})$  with  $M_{ij} \in \{0, 1\}$  iid Bernoulli with parameter d/n.

The non-symmetric observed matrix  $A \in M_n(\mathbb{R})$  is

$$A_{ij} = \frac{n}{d} P_{ij} M_{ij}.$$

The symmetric observed matrix  $H \in M_n(\mathbb{C})$  is

$$H = \frac{A + A^*}{2}.$$

Let  $P \in M_n(\mathbb{R})$  be a symmetric matrix.

Let  $M = (M_{ij}) \in M_n(\mathbb{R})$  with  $M_{ij} \in \{0, 1\}$  iid Bernoulli with parameter d/n.

The non-symmetric observed matrix  $A \in M_n(\mathbb{R})$  is

$$A_{ij} = \frac{n}{d} P_{ij} M_{ij}.$$

The symmetric observed matrix  $H \in M_n(\mathbb{C})$  is

$$H = \frac{A + A^*}{2}.$$

We have  $\mathbb{E}A = \mathbb{E}H = P$ , i.e. A and H are noisy unbiased observations of P.

We would like to compare the k-th largest eigenvalues of P and A or H and their corresponding eigenspaces

We would like to compare the k-th largest eigenvalues of P and A or H and their corresponding eigenspaces

It is smarter to consider the spectrum of A rather than the spectrum of H.

We would like to compare the k-th largest eigenvalues of P and A or H and their corresponding eigenspaces

It is smarter to consider the spectrum of A rather than the spectrum of H.

Benefits of asymmetry: in some situations, the spectrum of a matrix P is much less perturbed by a random asymmetric noise than by a random symmetric noise.

# DETECTION THRESHOLD

We set

$$Q_{ij} = n|P_{ij}|^2 \quad \text{and} \quad \rho = ||Q||.$$

# The detection threshold is defined as

$$\theta = \max\left(\sqrt{\frac{\rho}{d}}, \frac{L}{d}\right),$$

with

$$L = n \max_{i,j} |P_{ij}|.$$

## INCOHERENCE

We order the real eigenvalues of P,

$$|\mu_1| \ge \cdots \ge |\mu_{r_0}| > \theta \ge |\mu_{r_0+1}| \ge \cdots \ge |\mu_n|.$$

In an ON basis of eigenvectors of P, for all  $1 \leq k \leq r_0$ ,

$$\|\varphi_k\|_{\infty} = \max_i |\varphi_k(i)| \leqslant \frac{b}{\sqrt{n}}.$$

#### STABLE NUMERICAL RANK

The stable numerical rank is

$$r = \frac{\sum_k \mu_k^2}{\mu_1^2} \leqslant \operatorname{rank}(P).$$

In this talk, we assume that  $r, b, L, d, r_0$  are  $n^{o(1)}$ . All results are quantitative but will be stated in an asymptotic way.

# ESTIMATION OF EIGENVALUES

Recall:

$$A_{ij} = \frac{n}{d} P_{ij} M_{ij}$$
 and  $\theta = \max\left(\sqrt{\frac{\rho}{d}}, \frac{L}{d}\right).$ 

Eigenvalues of  ${\cal P}$  :

 $|\mu_1| \ge \cdots \ge |\mu_{r_0}| > \theta \ge |\mu_{r_0+1}| \ge \cdots \ge |\mu_n|.$ 

#### ESTIMATION OF EIGENVALUES

Recall:

$$A_{ij} = \frac{n}{d} P_{ij} M_{ij}$$
 and  $\theta = \max\left(\sqrt{\frac{\rho}{d}}, \frac{L}{d}\right).$ 

Eigenvalues of  ${\cal P}$  :

 $|\mu_1| \ge \cdots \ge |\mu_{r_0}| > \theta \ge |\mu_{r_0+1}| \ge \cdots \ge |\mu_n|.$ 

#### Theorem

With high probability, there exists an ordering of the eigenvalues  $\lambda_1, \ldots, \lambda_n$  of A such that

 $\max_{1\leqslant k\leqslant r_0}|\lambda_k-\mu_k|=o(1) \quad \text{ and } \quad \max_{r_0+1\leqslant k\leqslant n}|\lambda_k|\leqslant (1+o(1))\theta.$ 

# SIMULATION

For n = 2000 and  $P = 3\varphi_1\varphi_1^* + 2\varphi_2\varphi_2^* + \varphi_3\varphi_3^*$  with  $\varphi_k$  uniform.



# SIMULATION

For n = 2000 and  $P = 3\varphi_1\varphi_1^* + 2\varphi_2\varphi_2^* + \varphi_3\varphi_3^*$  with  $\varphi_k$  uniform.



Recall:

$$A_{ij} = \frac{n}{d} P_{ij} M_{ij}$$
 and  $\theta = \max\left(\sqrt{\frac{\rho}{d}}, \frac{L}{d}\right).$ 

We assume that the large eigenvalues of

$$P = \sum_{k} \mu_k \varphi_k \varphi_k^*$$

are well separated:

$$\left|1 - \frac{\mu_k}{\mu_l}\right| \geqslant \frac{\log d}{\log n}$$

for all  $1 \leq k \neq l \leq r_0$ .

Recall:

$$A_{ij} = \frac{n}{d} P_{ij} M_{ij}$$
 and  $\theta = \max\left(\sqrt{\frac{\rho}{d}}, \frac{L}{d}\right).$ 

We assume that the large eigenvalues of

$$P = \sum_{k} \mu_k \varphi_k \varphi_k^*$$

are well separated:

$$\left|1 - \frac{\mu_k}{\mu_l}\right| \geqslant \frac{\log d}{\log n} \quad \text{for all} \quad 1 \leqslant k \neq l \leqslant r_0.$$

# Theorem

Let  $\psi_k$  be a unit eigenvector associated to k-th eigenvalue of A. There exists  $\gamma_k > 0$  such that, with high probability, for  $1 \leq k \leq r_0$ ,  $|\langle \psi_k, \varphi_k \rangle| = \gamma_k + o(1).$ 

The asymptotic scalar product  $\gamma_k = |\langle \psi_k, \phi_k \rangle| + o(1)$  has an explicit formula:

$$\gamma_k = \frac{1}{\sqrt{\Gamma_{k,k}}}$$

with, for  $1 \leq k, l \leq r_0$ ,

$$\Gamma_{k,l} = \sum_{i=1}^{n} w_{k,l}(i)\varphi_k(i)\varphi_l(i),$$

and

$$w_{k,l}(i) = \sum_{j} \left( I - \frac{Q}{\mu_k \mu_l d} \right)_{i,j}^{-1}.$$

*Remark*:  $|\langle \psi_k, \psi_l \rangle| = |\Gamma_{k,l}| / \sqrt{\Gamma_{k,k}\Gamma_{l,l}} + o(1)$  is non-zero for  $k \neq l$  if **1** is not an eigenvector of Q.

RANK ONE PROJECTOR

If 
$$P = \varphi \varphi^*$$
, we find

$$\theta = \sqrt{\frac{n\sum_i |\varphi(i)|^4}{d}}$$

$$\gamma = \sqrt{1 - \frac{n\sum_i |\varphi(i)|^4}{d}}.$$

It is also possible to compute the scalar between the left  $\psi'_k$  and right  $\psi_k$  unit eigenvectors of the k-th eigenvalue of A:

$$\langle \psi'_k, \psi_k \rangle = \gamma_k^2 + o(1) = \frac{1}{\Gamma_{k,k}} + o(1).$$

We get an estimator

$$\hat{\varphi}_k = \frac{\psi_k + \psi'_k}{\|\psi_k + \psi'_k\|_2}$$

such that

$$|\langle \varphi_k, \hat{\varphi}_k \rangle| = \sqrt{\frac{2\gamma_k^2}{1+\gamma_k^2}} + o(1).$$

# SIMULATION

For n = 6000 and  $P = \varphi \varphi^*$  with  $\varphi$  uniform on the sphere.



IMPROVED ESTIMATION WITH NON-BACKTRACKING MATRICES

#### PUT SOME SYMMETRY BACK

We can improve the factor d in 2d in the detection threshold:

$$heta = \max\left(\sqrt{rac{
ho}{d}}, rac{L}{d}
ight).$$

We have not taken into account the information

$$P_{ij} = P_{ji}.$$

There is in fact an average of 2d observed entries per row.

#### NON-BACKTRACKING MATRIX

The set of symmetric observed entries is

 $E = \{(i, j) : (i, j) \text{ or } (j, i) \text{ is observed}\}.$ 

We have  $|E| \sim 2dn$ .

We consider the non-symmetric matrix  $B \in M_E(\mathbb{C})$  defined for all (i, j), (k, l) in E by

$$B_{(i,j),(k,l)} = \frac{nP_{kl}}{2d}\mathbf{1}(j=k, l\neq i)$$

$$e \qquad f$$

$$i \qquad j=k \qquad l\neq i$$

٠

#### NON-BACKTRACKING MATRIX

A vector  $\varphi \in \mathbb{C}^n$  is lifted in  $\mathbb{C}^E$  as

 $\varphi^+(i,j) = \varphi(j).$ 

#### Theorem

The preceding results on A are true for the matrix B (with minor extra changes) with d replaced by 2d.

# SIMULATION

For n = 5000 and  $P = \varphi \varphi^*$  with  $\varphi$  uniform on the sphere.



# THE DETECTION THRESHOLD

$$\theta = \max\left(\sqrt{\frac{\rho}{d}}, \frac{L}{d}\right).$$

#### LIFT OF A MATRIX

Fix  $j_0 \in [n] = \{1, \ldots, n\}$ . Let V be the set of finite integer sequences in  $[n], (j_0, j_1, \ldots, j_k)$  starting with  $j_0$ .

We build an infinite matrix  $\mathcal{P} = (\mathcal{P}_{uv})_{u,v \in V}$  by setting  $u = (j_0, \ldots, j_k) \in V$  and  $j \in [n]$ 

$$\mathcal{P}_{u,(u,j)} = P_{j_k,j}.$$



This defines a non-symmetric bounded operator on  $\ell^2(V)$  build on an infinite *n*-ary tree.

LIFT OF A MATRIX

If  $P\varphi = \mu\varphi$  then  $\Phi$  defined on V as

 $\Phi(j_0\cdots j_k)=\varphi(j_k)$ 

satisfies

 $\mathcal{P}\Phi = \mu\Phi.$ 



The function  $\Phi$  is not in  $\ell^2(V)$ .

We keep each edge with probability d/n.



We denote by  $\mathcal{P}_{\text{perc}}$  the corresponding operator and set

$$\mathcal{A} = \frac{n}{d} \mathcal{P}_{\text{perc}}.$$

The operator  $\mathcal{A}$  is a local approximation of the matrix  $A = (n/d)P \odot M$ .





Since  $\mathcal{P}\Phi = \mu\Phi$ , for all  $v \in V$ , the process in  $t \in \mathbb{N}$ ,

 $\Psi_t(v) = \mu^{-t}(\mathcal{A}^t \Phi)(v)$ 

is a discrete martingale for the filtration of the successive generations in the tree.

The bracket of the martingale can be computed and we find:

$$\mathbb{E}|\Psi_{t+1}(v) - \Psi_t(v)|^2 = \frac{Q^t(\varphi^2)(v)}{(\mu^2 d)^t}.$$

Recall:  $Q_{ij} = n|P_{ij}|^2$  and  $\rho = ||Q||$ .

Hence  $\Psi_t(v) = \mu^{-t}(\mathcal{A}^t \Phi)(v)$  converges a.s. and in  $L^2$  toward  $\Psi(v)$  if  $|\mu| > \sqrt{\frac{\rho}{d}}.$ 

This is called the Kesten-Stigum threshold.

If  $|\mu| > \sqrt{\rho/d}$ , then  $\Psi_t(v) = \mu^{-t}(\mathcal{A}^t \Phi)(v)$  converges a.s. and in  $L^2$  toward  $\Psi(v)$ .

Since

$$\mathcal{A}\Psi_t = \mu \Psi_{t+1},$$

we can define a.s. a random eigenwave  $\Psi$  on V which satisfies

 $\mathcal{A}\Psi=\mu\Psi.$ 

#### BACK TO FINITE DIMENSION

This analysis and concentration inequalities allow to show that if  $t \gg 1$  but not too large,

$$||A^{t+1}\varphi_k - \mu_k A^t \varphi_k||_2 = o(||A^t \varphi_k||_2).$$

Similarly  $\varphi_k^* A^t$  is an approximate left eigenvector.

We decompose  $A^t$  in

$$A^{t} = \sum_{k=1}^{r_0} \mu_k^{t} u_k v_k^* + R_t,$$

with  $u_k = A^t \varphi_k / \mu_k^t$  and  $v_k = (A^t)^* \varphi_k / \mu_k^t$ . We have

$$\langle u_k, v_l \rangle = \delta_{kl} + o(1).$$

#### PROOF STRATEGY

For  $t = c \ln n / \log d$  well chosen,

$$A^{t} = \sum_{k=1}^{r_0} \mu_k^{t} u_k v_k^* + R_t.$$

- \* Compute the inner products between these  $r_0^2$  vectors;
- $\star$  Show that the Gram matrix is well-conditioned;
- \* Show that  $||R_t|| \leq (\log n)^c \theta^t$ ;
- ★ Use an ad-hoc spectral perturbation theorem of a non-symmetric matrix of Bauer-Fike type.

B-Lelarge-Massoulié 18.

# RECTANGULAR MATRICES

#### LINEARIZATION TRICK

If  $P \in M_{m,n}(\mathbb{C})$ , the matrix

$$\widetilde{P} = \begin{pmatrix} 0 & P \\ P^* & 0 \end{pmatrix}$$

is of size  $(m+n) \times (m+n)$  and is Hermitian.

The singular value decomposition of  $P = \sum_k s_k u_k v_k^*$  is equivalent to the diagonalization of  $\widetilde{P}$ :

$$\widetilde{P} = \sum_{k} s_k w_k^+ (w_k^+)^* - s_k w_k^- (w_k^-)^*,$$

with  $w_k^{\pm} = (u_k, \pm v_k)' / \sqrt{2}$ .

### A RANDOMIZED ASYMMETRIC SVD

Recall

$$P = \sum_{k} s_k u_k v_k^*.$$

Consider  $Z = (Z_{ij}) \in M_{m,n}(\mathbb{R})$  with iid  $\{0, 1\}$ -Bernoulli entries with parameter 1/2 and define

$$P_1 P_2^*$$
 with  $P_1 = P \odot Z$ ,  $P_2 = P - P_1$ .

The k-th largest eigenvalue, say  $\lambda_k$ , of  $P_1 P_2^*$  is a proxy for  $s_k^2/4$ .

The average of the left and right eigenvectors associated to  $\lambda_k$  is a proxy for the left singular vector  $u_k$ .

#### MATRIX COMPLETION

Let  $M = (M_{ij}) \in M_{m,n}(\mathbb{R})$  with iid  $\{0, 1\}$ -Bernoulli entries with parameter d/n.

The observed matrix is

$$A = \frac{n}{d} P \odot M.$$

We perform the randomized asymmetric SVD on A.

At a higher computational cost, we may also consider the non-backtracking matrix associated to the linearized matrix  $\widetilde{A}$ .

In either case, if  $n \simeq m$ , we have explicit detection thresholds and formulas for the asymptotic inner products.

### MATRIX COMPLETION

## Recall

$$P = \sum_{k} s_k u_k v_k^*.$$

Once we have estimators  $\hat{u}_k$ ,  $\hat{v}_k$  of  $u_k$  and  $v_k$ , it is possible to design an estimator of P:

$$\hat{P} = \sum_{k=1}^{r_0} x_k \hat{u}_k \hat{v}_k^*$$

for some vector  $x = (x_k) \in \mathbb{R}^{r_0}$  which asymptotically minimizes

 $\|\hat{P} - P\|_F$ 

and compute an explicit asymptotic formula for

$$\mathrm{MSE}(\hat{P}) = \|\hat{P} - P\|_F^2$$

Nadakuditi 14.

### SIMULATION

We take d = 9.7, (m, n) = (2000, 3000) and  $P = uv^*$  with u, v independent standard Gaussian vectors.



# CONCLUDING WORDS

# CONCLUSION

Spectral analysis methods on random non-symmetric matrices can be very efficient, *Chen-Cheng-Fan 18.* 

Numerous possible extensions, for example include some extra noise, or models where the probability of observing an entry depends on the entry, *Stephan-Massoulié 20*.

There is nowadays a lot of activities on tensor completion.

THANK YOU FOR YOUR ATTENTION!